

MA4270 Data Modelling and Computation

Final Examination Helpsheet

AY2023/24 Semester 2 · Prepared by Tian Xiao @snoidetx

1 Perceptron

Classification Problems: To learn a classifier f_{θ} that classifies labels.

- Dataset: $\mathcal{D} = \{(\mathbf{x}_t, y_t)\}_{t=1}^n$ where $\mathbf{x}_t \in \mathbb{R}^d$ and $y_t \in \{-1, +1\}$.
- Classifier: $f_{\theta} : \mathbb{R}^d \rightarrow \{-1, +1\}$.
 - Linear classifier: $f_{\theta} = \text{sign}(\theta^\top \mathbf{x})$.
- Training error: $\hat{E}(\theta) = \frac{1}{n} \sum_{t=1}^n \text{Loss}(y_t, f_{\theta}(\mathbf{x}_t))$.
 - 0-1 loss: $\text{Loss}(y, \hat{y}) = \ell(y, \hat{y}) = \mathbf{1}\{\hat{y} \neq y\} = \begin{cases} 1 & \hat{y} \neq y \\ 0 & \text{otherwise} \end{cases}$.
 - Linear separable: $\exists \theta [\hat{E}(\theta) = 0]$.

The Perceptron Algorithm:

- Initialize $\theta^{(0)}$ to some value (e.g., $\mathbf{0}$), and initialize index k to 0.
 - Repeatedly perform the following:
 - Select the next example (\mathbf{x}_t, y_t) from the training set and check whether $\theta^{(k)}$ classifies it correctly.
 - If it is incorrect (i.e., $y_t (\theta^{(k)})^\top \mathbf{x}_t < 0$), set $\theta^{(k+1)} \leftarrow \theta^{(k)} + y_t \mathbf{x}_t$ and increment $k \leftarrow k + 1$.
- Assumptions:**
- Inputs are bounded: $\exists R \in (0, \infty) \forall \mathbf{x}_t \in \mathcal{D} [\|\mathbf{x}_t\| \leq R]$.
 - Linearly separable: $\exists \theta^* \exists \gamma > 0 [\min_t y_t (\theta^*)^\top \mathbf{x}_t \geq \gamma]$.
- Convergence.** Under the initial vector $\theta^{(0)} = \mathbf{0}$, for any dataset \mathcal{D} satisfying the above assumptions, the perceptron algorithm produces a vector $\theta^{(k)}$ classifying every example correctly after at most $k_{\max} = \frac{R^2 \|\theta^*\|^2}{\gamma^2}$ mistakes (and hence update steps).

Proof. Let $R = \max \|\mathbf{x}_t\|$, $\gamma = \min_t y_t (\theta^*)^\top \mathbf{x}_t$ for $t = 1, 2, \dots, n$.

- $(\theta^*)^\top \theta^{(k)} = (\theta^*)^\top (\theta^{(k-1)} + y_t \mathbf{x}_t) \geq (\theta^*)^\top \theta^{(k-1)} + \gamma$. So $(\theta^*)^\top \theta^{(k)} \geq k\gamma$.
- $\|\theta^{(k)}\|^2 = \|\theta^{(k-1)}\|^2 + 2(\theta^{(k-1)}, y_t \mathbf{x}_t) + \|\mathbf{x}_t\|^2 \leq \|\theta^{(k-1)}\|^2 + \|\mathbf{x}_t\|^2$. So $\|\theta^{(k)}\|^2 \leq kR^2$.
- By Cauchy-Schwarz inequality $\langle \mathbf{v}, \mathbf{w} \rangle \leq \|\mathbf{v}\| \cdot \|\mathbf{w}\|$, we have $1 \geq \frac{\langle \theta^{(k)}, \theta^* \rangle}{\|\theta^{(k)}\| \cdot \|\theta^*\|} \geq \frac{k\gamma}{\sqrt{kR^2} \|\theta^*\|}$, hence $k \leq \frac{R^2 \|\theta^*\|^2}{\gamma^2}$.

- Margin: Let $\gamma = \min_{t=1,2,\dots,n} y_t \theta^{(k)} \mathbf{x}_t$. The quantity $\gamma_{\text{geom}} = \frac{\gamma}{\|\theta\|}$ is the smallest distance from any example \mathbf{x}_t to the decision boundary specified by θ .

2 Support Vector Machine (SVM)

Maximum Margin Classifier: $\min_{\theta} \frac{1}{2} \|\theta\|^2$ s.t. $\forall t [y_t \theta^\top \mathbf{x}_t \geq 1]$ (unique)

- SVM with offset: $\min_{\theta, \theta_0} \frac{1}{2} \|\theta\|^2$ s.t. $\forall t [y_t (\theta^\top \mathbf{x}_t + \theta_0) \geq 1]$.
 - Support vectors: On margin ($y_t (\theta^\top \mathbf{x}_t + \theta_0) = 1$).
- Soft-margin SVM: $\min_{\theta, \theta_0, \zeta} \frac{1}{2} \|\theta\|^2 + C \sum_{t=1}^n \zeta_t$ s.t. $\forall t [y_t (\theta^\top \mathbf{x}_t + \theta_0) \geq 1 - \zeta_t]$.
 - $\zeta = (\zeta_1, \zeta_2, \dots, \zeta_n) \geq \mathbf{0}$ is called *slack variables*.
 - Support vectors: On margin/within margin/misclassified.
- Hinge-loss formulation: $\min_{\theta, \theta_0} \frac{1}{2} \|\theta\|^2 + C \sum_{t=1}^n [1 - y_t (\theta^\top \mathbf{x}_t + \theta_0)]_+$.
 - Hinge loss: $z \rightarrow [1 - z]_+ = \max\{0, 1 - z\}$.
 - Interpretation: Total hinge loss with regularization term $\frac{1}{2} \|\theta\|^2$.

3 Logistic Regression

Logistic Likelihood Model: $\Pr(y | \mathbf{x}) = \frac{1}{1 + \exp(-y(\theta^\top \mathbf{x} + \theta_0))}$.

- $g(z) = \frac{1}{1+z} \in (0, 1)$ assigns *likelihood* to points.
 - Scaling the dataset by $c > 1$ pushes prediction closer to 0 or 1.
 - Linear classifier chooses the label that is more likely under the logistic model.
 - Log-odds log $\log \frac{\Pr(y=1|\mathbf{x})}{\Pr(y=-1|\mathbf{x})}$ is a linear function $\langle \theta, \mathbf{x} \rangle + \theta_0$ of inputs.
- Maximum likelihood estimate (MLE) of parameters:

$$\begin{aligned} (\theta, \theta_0) &= \arg \max_{\theta, \theta_0} \prod_{t=1}^n \Pr(y_t | \mathbf{x}_t; \theta, \theta_0) \quad (\text{likelihood}) \\ &= \arg \max_{\theta, \theta_0} \prod_{t=1}^n \frac{1}{1 + \exp(-y_t (\theta^\top \mathbf{x}_t + \theta_0))} \quad (\text{likelihood}) \\ &= \arg \max_{\theta, \theta_0} \sum_{t=1}^n \log \frac{1}{1 + \exp(-y_t (\theta^\top \mathbf{x}_t + \theta_0))} \quad (\text{log-likelihood}) \\ &= \arg \min_{\theta, \theta_0} \sum_{t=1}^n \log (1 + \exp(-y_t (\theta^\top \mathbf{x}_t + \theta_0))). \end{aligned}$$

- Regularization: $\min_{\theta, \theta_0} \frac{1}{2} \|\theta\|^2 + C \sum_{t=1}^n \log (1 + \exp(-y_t (\theta^\top \mathbf{x}_t + \theta_0)))$.

▷ Logistic loss: $z \rightarrow \log(1 + e^{-z})$.

▷ Interpretation: Total logistic loss with regularization term $\frac{1}{2} \|\theta\|^2$.

- Softmax function (multiclass): $\Pr(y = c | \mathbf{x}) = \frac{\exp(\theta_c^\top \mathbf{x} + \theta_{0,c})}{\sum_{c'=1}^M \exp(\theta_{c'}^\top \mathbf{x} + \theta_{0,c'})}$.
 - When $M = 2$, we recover logistic model by setting $(\theta_c, \theta_{0,c}) = (\mathbf{0}, 0)$ for one of the two classes.

4 Linear Regression

Linear Predictor: $\hat{y} = \theta^\top \mathbf{x} + \theta_0$.

- Matrix form: $\hat{\mathbf{y}} = \mathbf{X}\Theta$, where $\mathbf{X} = \begin{bmatrix} \mathbf{x}_1^\top & 1 \\ \vdots & \vdots \\ \mathbf{x}_n^\top & 1 \end{bmatrix}$ and $\Theta = \begin{bmatrix} \theta \\ \theta_0 \end{bmatrix}$.
- Least squares estimate (LSE): $\hat{\Theta} = (\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top \mathbf{y}$.
 - Unique solution if $\mathbf{X}^\top \mathbf{X}$ is invertible.
- Gaussian model: $y_t = (\theta^*)^\top \mathbf{x}_t + \theta_0^* + z_t$, where $z_t \sim \mathcal{N}(0, \sigma^2)$.
 - Gaussian PDF: $\mathcal{N}(z; \mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(z-\mu)^2}{2\sigma^2}\right)$.
 - $\Pr(y | \mathbf{x}) = \mathcal{N}(y; (\theta^*)^\top \mathbf{x} + \theta_0^*, \sigma^2)$.
 - Log-likelihood:

$$\log \prod_{t=1}^n \Pr(y_t | \mathbf{x}_t) = \text{const.} - \frac{n}{2} \log \sigma^2 - \frac{1}{2\sigma^2} \sum_{t=1}^n (y_t - \theta^\top \mathbf{x}_t - \theta_0)^2$$
 - MLE of θ and θ_0 : $(\hat{\theta}, \hat{\theta}_0) = \arg \min_{\theta, \theta_0} \sum_{t=1}^n (y_t - \theta^\top \mathbf{x}_t - \theta_0)^2$.
 - σ^2 is assumed to be known.
 - σ^2 MLE of σ^2 : $\hat{\sigma}^2 = \frac{1}{n} \sum_{t=1}^n (y_t - \hat{\theta}^\top \mathbf{x}_t - \hat{\theta}_0)^2$.
- Gaussian model in matrix form: $\mathbf{y} = \mathbf{X}\Theta^* + \mathbf{z}$.
 - LSE: $\hat{\Theta} = \Theta^* + (\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top \mathbf{z}$.
 - No bias: $\mathbb{E}[\hat{\Theta}] = \Theta^*$.
 - Covariance: $\text{Cov}[\hat{\Theta}] = \sigma^2 (\mathbf{X}^\top \mathbf{X})^{-1}$.
- Ridge regression: $(\hat{\theta}, \hat{\theta}_0) = \arg \min_{\theta, \theta_0} \sum_{t=1}^n (y_t - \theta^\top \mathbf{x}_t - \theta_0)^2 + \lambda \sum_{j=1}^d \theta_j^2$.
 - Closed-form solution (w/o offset): $\hat{\theta} = (\mathbf{X}^\top \mathbf{X} + \lambda \mathbf{I})^{-1} \mathbf{X}^\top \mathbf{y}$.
 - $\mathbf{X}^\top \mathbf{X} + \lambda \mathbf{I}$ is always invertible when $\lambda > 0$.
 - Assuming no offset θ_0 :
 - Bias: $\mathbb{E}[\hat{\theta}] - \Theta^* = -\lambda (\mathbf{X}^\top \mathbf{X} + \lambda \mathbf{I})^{-1} \Theta^*$.
 - Covariance: $\sigma^2 ((\mathbf{X}^\top \mathbf{X} + \lambda \mathbf{I})^{-1} - \lambda (\mathbf{X}^\top \mathbf{X} + \lambda \mathbf{I})^{-2})$.

Bias-Variance Tradeoff: Decomposition of MSE:

$$\mathbb{E}[\|\hat{\Theta} - \Theta^*\|^2] = \underbrace{\mathbb{E}[\|\hat{\Theta}\|^2]}_{\text{bias}} - 2\mathbb{E}[\langle \hat{\Theta}, \Theta^* \rangle] + \underbrace{\mathbb{E}[\|\Theta^*\|^2]}_{\text{variance}}$$

Proof. Let $\mu = \mathbb{E}[\hat{\Theta}]$.

- bias = $\|\mu\|^2 - 2\langle \mu, \Theta^* \rangle + \|\Theta^*\|^2$.
- variance = $\mathbb{E}[\|\hat{\Theta}\|^2] - 2\langle \hat{\Theta}, \mu \rangle + \|\mu\|^2 = \mathbb{E}[\|\hat{\Theta}\|^2] - 2\langle \mathbb{E}[\hat{\Theta}], \mu \rangle + \|\mu\|^2 = \mathbb{E}[\|\hat{\Theta}\|^2] - \|\mu\|^2$.
- bias + variance = $\mathbb{E}[\|\hat{\Theta}\|^2] - 2\langle \mu, \Theta^* \rangle + \|\Theta^*\|^2 = \text{LHS}$.

5 Kernel Method

Kernel: A measure of similarity.

- Kernel matrix: A function $k : \mathbb{R}^d \times \mathbb{R}^d \rightarrow \mathbb{R}$ is said to be a *positive semidefinite* (PSD) kernel if
 - it is symmetric, i.e., $k(\mathbf{x}, \mathbf{x}') = k(\mathbf{x}', \mathbf{x})$;
 - the following *kernel matrix* is always PSD:

$$\mathbf{K} = \begin{bmatrix} k(\mathbf{x}_1, \mathbf{x}_1) & \cdots & k(\mathbf{x}_1, \mathbf{x}_m) \\ \vdots & \ddots & \vdots \\ k(\mathbf{x}_m, \mathbf{x}_1) & \cdots & k(\mathbf{x}_m, \mathbf{x}_m) \end{bmatrix}.$$
- Polynomial kernel: $k(\mathbf{x}, \mathbf{x}') = (1 + \langle \mathbf{x}, \mathbf{x}' \rangle)^p$.
- RBF kernel: $k(\mathbf{x}, \mathbf{x}') = \exp(-\frac{1}{2} \|\mathbf{x} - \mathbf{x}'\|^2)$.
- k is a PSD kernel iff it equals an inner product $\langle \phi(\mathbf{x}), \phi(\mathbf{x}') \rangle$ for some (possibly infinite dimensional) mapping ϕ .
- Construction: If k_1, k_2 are kernels, then the following are kernels:
 - $k(\mathbf{x}, \mathbf{x}') = f(\mathbf{x}) k_1(\mathbf{x}, \mathbf{x}') f(\mathbf{x}')$ for some function f ;
 - $k(\mathbf{x}, \mathbf{x}') = k_1(\mathbf{x}, \mathbf{x}') + k_2(\mathbf{x}, \mathbf{x}')$;
 - $k(\mathbf{x}, \mathbf{x}') = k_1(\mathbf{x}, \mathbf{x}') k_2(\mathbf{x}, \mathbf{x}')$.

Kernel Trick: Replace $\langle \mathbf{x}, \mathbf{x}' \rangle$ by $k(\mathbf{x}, \mathbf{x}')$.

- Possible when algorithm depends on only inputs' inner products.
- Dual → Kernel SVM: $\alpha > 0$ support vectors; $\alpha = C$ violation.

$$\begin{array}{ll} (\text{P}) \min_{\theta, \theta_0, \zeta} & \frac{1}{2} \|\theta\|^2 + C \sum_{t=1}^n \zeta_t \\ \text{s.t.} & y_t (\theta^\top \mathbf{x}_t + \theta_0) \geq 1 - \zeta_t; \\ & \zeta_t \geq 0, \forall t. \end{array} \quad \begin{array}{ll} (\text{D}) \max_{\alpha} & \sum_{t=1}^n \alpha_t - \frac{1}{2} \sum_{s=1}^n \sum_{t=1}^n \alpha_s \alpha_t y_s y_t k(\mathbf{x}_s, \mathbf{x}_t) \\ \text{s.t.} & \alpha_t \in [0, C], \forall t; \\ & \sum_{t=1}^n \alpha_t y_t = 0. \end{array}$$

6 Gradient-Based Optimization

Gradient Descent: W.r.t. $f(\mathbf{x})$, $\mathbf{x}_{\text{next}} = \mathbf{x} - \eta \cdot \nabla f(\mathbf{x})$.

- Stochastic gradient descent (SGD): $\mathbf{x}_{\text{next}} = \mathbf{x} - \eta \cdot \nabla f_i(\mathbf{x})$.
- Mini-batch SGD: $\mathbf{x}_{\text{next}} = \mathbf{x} - \eta \cdot \frac{1}{|\mathcal{I}|} \sum_{i \in \mathcal{I}} \nabla f_i(\mathbf{x})$.

Subgradient-Based Optimization: Non-differentiable convex functions.

- Subgradient: $\partial f(\mathbf{x}) = \{\mathbf{g} \in \mathbb{R}^d : f(\mathbf{x}') \geq f(\mathbf{x}) + \langle \mathbf{g}, \mathbf{x}' - \mathbf{x} \rangle, \forall \mathbf{x}'\}$.
 ▷ $f(x) = |x|$: $\partial = \{1\}$ at $x > 0$; $\{-1\}$ at $x < 0$; $[-1, 1]$ at $x = 0$.
- Subgradient method: $\mathbf{x}_{\text{next}} = \mathbf{x} - \eta \cdot \mathbf{g}$, $\mathbf{g} \in \partial f(\mathbf{x})$.
 ▷ **Convergence.** Assume that f is convex, minimizer \mathbf{x}^* exists, Lipschitz condition ($\|\mathbf{g}\| \leq M, \forall \mathbf{g}, \mathbf{x}$) holds, initialization $\mathbf{x}^{(1)}$ satisfies $\|\mathbf{x}^{(1)} - \mathbf{x}^*\| \leq R$ for some finite R . Using subgradient method with any sequence of step sizes $\{\eta_k\}_{k=1}^\infty$ satisfying $\lim_{k \rightarrow \infty} \eta_k = 0$ and $\sum_{k=1}^\infty \eta_k = \infty$, we have as $k \rightarrow \infty$

$$\min_{k=1, \dots, K} f(\mathbf{x}^{(k)}) \rightarrow f(\mathbf{x}^*).$$

* Choosing $\eta_k = \frac{\eta_0}{\sqrt{k}}$, we yield a convergence rate of $O(\frac{\log k}{\sqrt{k}})$.

Proof. How close the $(k+1)$ -th iterate is to \mathbf{x}^* ?

$$\begin{aligned} \frac{1}{2} \|\mathbf{x}^{(k+1)} - \mathbf{x}^*\|^2 &= \frac{1}{2} \|\mathbf{x}^{(k)} - \eta_k \mathbf{g}^{(k)} - \mathbf{x}^*\|^2 \\ &= \frac{1}{2} \|\mathbf{x}^{(k)} - \mathbf{x}^*\|^2 - \eta_k \mathbf{g}^{(k) \top} (\mathbf{x}^{(k)} - \mathbf{x}^*) + \frac{\eta_k^2}{2} \|\mathbf{g}^{(k)}\|^2 \\ &\leq \frac{1}{2} \|\mathbf{x}^{(k)} - \mathbf{x}^*\|^2 - \eta_k (f(\mathbf{x}^{(k)}) - f(\mathbf{x}^*)) + \frac{\eta_k^2}{2} \|\mathbf{g}^{(k)}\|^2. \end{aligned}$$

Rearranging and summing from $k=1$ to K :

$$\begin{aligned} \sum_{k=1}^K \eta_k (f(\mathbf{x}^{(k)}) - f(\mathbf{x}^*)) &\leq \frac{1}{2} \|\mathbf{x}^{(1)} - \mathbf{x}^*\|^2 - \frac{1}{2} \|\mathbf{x}^{(K+1)} - \mathbf{x}^*\|^2 + \sum_{k=1}^K \frac{\eta_k^2}{2} \|\mathbf{g}^{(k)}\|^2 \\ &\leq \frac{1}{2} \|\mathbf{x}^{(1)} - \mathbf{x}^*\|^2 + \sum_{k=1}^K \frac{\eta_k^2}{2} \|\mathbf{g}^{(k)}\|^2 \\ &\leq \frac{1}{2} R^2 + \frac{1}{2} M^2 \sum_{k=1}^K \eta_k^2. \\ \min_{k=1, \dots, K} f(\mathbf{x}^{(k)}) - f(\mathbf{x}^*) &\leq \frac{\frac{1}{2} R^2 + \frac{1}{2} M^2 \sum_{k=1}^K \eta_k^2}{\sum_{k=1}^K \eta_k} \rightarrow 0 \text{ as } K \rightarrow \infty. \end{aligned}$$

Projected Gradient-Based Optimization: Constrained problems.

- Projected gradient descent: $\mathbf{x}_{\text{next}} = \Pi_{\mathcal{C}}(\mathbf{x} - \eta \cdot \mathbf{g})$.
 ▷ Projected to the closest point in feasible set \mathcal{C} .

7 Boosting

Decision Stump: $h(\mathbf{x}, \theta) = h(\mathbf{x}, \{s, k, \theta_0\}) = \text{sign}(s(x_k - \theta_0))$.

- Choose $s \in \{1, -1\}$ s.t. h is $\frac{1}{2}$ correct.

AdaBoost: Weighted aggregation of simple models (decision stumps).

- Exponential loss: $(y, f(\mathbf{x})) \rightarrow \exp(-yf(\mathbf{x}))$.

① Initialize $w_0(t) = 1/n$ for $t = 1, \dots, n$.
② For $m = 1, \dots, M$ do
▷ Choose the next base learner $h(\cdot, \hat{\theta}_m)$ as
$\hat{\theta}_m = \arg \min_{\theta} \sum_{t: y_t \neq h(\mathbf{x}_t, \theta)} w_{m-1}(t).$
▷ Set $\hat{\alpha}_m = \frac{1}{2} \log \frac{1-\hat{\epsilon}_m}{\hat{\epsilon}_m}$, where $\hat{\epsilon}_m = \sum_{t: y_t \neq h(\mathbf{x}_t, \hat{\theta}_m)} w_{m-1}(t)$.
▷ Update the weights and normalize by Z_m :
$w_m(t) = \frac{1}{Z_m} w_{m-1}(t) e^{-y_t h(\mathbf{x}_t, \hat{\theta}_m) \hat{\alpha}_m},$
$Z_m = \sum_{t=1}^n w_{m-1}(t) e^{-y_t h(\mathbf{x}_t, \hat{\theta}_m) \hat{\alpha}_m}.$
③ Output $f_M(\mathbf{x}) = \sum_{m=1}^M \hat{\alpha}_m h(\mathbf{x}, \hat{\theta}_m)$ w.r.t. classifier $\text{sign}(f_M(\mathbf{x}))$.

- Sum of weight w of wrongly classified examples is $1/2$.
- **Convergence.** After M iterations, the training error satisfies

$$\frac{1}{n} \sum_{t=1}^n \mathbf{1}\{y_t f_M(\mathbf{x}_t) \leq 0\} \leq \exp\left(-2 \sum_{m=1}^M \left(\frac{1}{2} - \hat{\epsilon}_m\right)^2\right).$$

In particular, if $\hat{\epsilon}_m \leq \frac{1}{2} - \gamma$ for all m and some $\gamma > 0$, then

$$\frac{1}{n} \sum_{t=1}^n \mathbf{1}\{y_t f_M(\mathbf{x}_t) \leq 0\} \leq \exp(-2M\gamma^2).$$

Proof. The 0-1 loss is bounded by exponential loss:

$$\begin{aligned} \frac{1}{n} \sum_{t=1}^n \mathbf{1}\{y_t f_M(\mathbf{x}_t) \leq 0\} &\leq \frac{1}{n} \sum_{t=1}^n e^{-y_t f_M(\mathbf{x}_t)} = \prod_{m=1}^M Z_m. \\ Z_m &= \sum_{t: y_t \neq h(\mathbf{x}_t, \theta)} e^{\hat{\alpha}_m w_{m-1}(t)} + \sum_{t: y_t = h(\mathbf{x}_t, \theta)} e^{-\hat{\alpha}_m w_{m-1}(t)} \\ &= e^{\hat{\alpha}_m \hat{\epsilon}_m} + e^{-\hat{\alpha}_m (1 - \hat{\epsilon}_m)} \\ &= \sqrt{\frac{1 - \hat{\epsilon}_m}{\hat{\epsilon}_m}} \hat{\epsilon}_m + \sqrt{\frac{\hat{\epsilon}_m}{1 - \hat{\epsilon}_m}} (1 - \hat{\epsilon}_m) \\ &= 2\sqrt{\hat{\epsilon}_m(1 - \hat{\epsilon}_m)} = \sqrt{1 - (1 - 2\hat{\epsilon}_m)^2} \leq e^{-\frac{1}{2}(1 - 2\hat{\epsilon}_m)^2}. \end{aligned}$$

Combine the two equations above and prove our theorem:

$$\frac{1}{n} \sum_{t=1}^n \mathbf{1}\{y_t f_M(\mathbf{x}_t) \leq 0\} \leq \prod_{m=1}^M Z_m \leq \prod_{m=1}^M e^{-\frac{1}{2}(1 - 2\hat{\epsilon}_m)^2} = e^{-2 \sum_{m=1}^M (\frac{1}{2} - \hat{\epsilon}_m)^2}.$$

8 Statistical Learning Theory

Hoeffding's Inequality: Let $Z = X_1 + \dots + X_n$, where $X_i \in [a_i, b_i]$:

$$\Pr\left[\frac{1}{n} |Z - \mathbb{E}[Z]| > \epsilon\right] \leq 2 \exp\left(-\frac{2n\epsilon^2}{\frac{1}{n} \sum_{i=1}^n (b_i - a_i)^2}\right).$$

Empirical Risk Minimization:

- True risk: $R(f) = \mathbb{E}[\ell(y, f(\mathbf{x}))]$.
- Empirical risk: $R_n(f) = \frac{1}{n} \sum_{i=1}^n \ell(y_i, f(\mathbf{x}_i))$.
- Test err $R(f) = \text{train err } R_n(f) + \text{generalization err } (R(f) - R_n(f))$.

PAC Learning: Given a loss function $\ell(\cdot, \cdot)$, a function class \mathcal{F} is said to be a *PAC-learnable* if there exists an algorithm $\mathcal{A}(\mathcal{D}_n)$ and a function $\bar{n}(\epsilon, \delta)$ such that for any distribution P_{XY} used to generate \mathcal{D}_n and any $\epsilon, \delta \in (0, 1)$, if $n > \bar{n}(\epsilon, \delta)$, the following holds with probability at least $1 - \delta$:

$$R(\hat{f}) \leq \min_{f \in \mathcal{F}} R(f) + \epsilon.$$

- ① The probability $1 - \delta$ corresponds to *probably correct*.
- ② The error ϵ corresponds to *approximately correct*.
- ③ The function \bar{n} is called the *sample complexity*.

- **Thm.** For any bounded loss function $\ell(y, f(\mathbf{x})) \in [0, 1]$, any finite function class \mathcal{F} is PAC-learnable with sample complexity

$$\bar{n}(\epsilon, \delta) = \frac{2}{\epsilon^2} \log \frac{2|\mathcal{F}|}{\delta}.$$

Proof. By Hoeffding's inequality,

$$\Pr[|R(f) - R_n(f)| \geq \epsilon_0] \leq 2e^{-2n\epsilon_0^2}.$$

Apply the union bound $\Pr[A_1 \cup \dots \cup A_m] \leq \sum_{i=1}^m \Pr[A_i]$:

$$\Pr\left[\bigcup_{f \in \mathcal{F}} \{|R(f) - R_n(f)| \geq \epsilon_0\}\right] \leq 2|\mathcal{F}| e^{-2n\epsilon_0^2}.$$

Setting RHS as δ , we get $n = \frac{1}{2\epsilon_0^2} \log \frac{2|\mathcal{F}|}{\delta}$, or $\epsilon_0 = \sqrt{\frac{1}{2n} \log \frac{2|\mathcal{F}|}{\delta}}$.

Let $f^* = \arg \min_{f \in \mathcal{F}} R(f)$. With probability $1 - \delta$,

$$\begin{aligned} R(f_{\text{erm}}) - R(f^*) &= R(f_{\text{erm}}) - R_n(f_{\text{erm}}) + R_n(f_{\text{erm}}) - R_n(f^*) + \\ &\quad R_n(f^*) - R(f^*) \\ &\leq \epsilon_0 + 0 + \epsilon_0 = 2\epsilon_0. \end{aligned}$$

Setting $\epsilon_0 = \epsilon/2$ and $\bar{n}(\epsilon, \delta) = \frac{2}{\epsilon^2} \log \frac{2|\mathcal{F}|}{\delta}$, we have proven.

- ▷ **Col.** With probability at least $1 - \delta$, for all $f \in \mathcal{F}$ we have:

$$|R(f) - R_n(f)| \leq \frac{1}{2n} \log \frac{2|\mathcal{F}|}{\delta}.$$

- ▷ **Col.** With probability at least $1 - \delta$,

$$R(f_{\text{erm}}) - \min_{f \in \mathcal{F}} R(f) \leq \sqrt{\frac{2}{n} \log \frac{2|\mathcal{F}|}{\delta}}.$$

Infinite Hypothesis Class:

- Growth function: Given any n unlabelled data, how many different assignments of labels can functions in \mathcal{F} make?

$$S_n(\mathcal{F}) = \sup_{\mathbf{x}_1, \dots, \mathbf{x}_n} |\{(f(\mathbf{x}_1), \dots, f(\mathbf{x}_n)) : f \in \mathcal{F}\}|.$$

- VC-dimension: Largest k such that $S_k(\mathcal{F}) = 2^k$ (can be ∞).
 ▷ If \mathcal{F} can provide 2^k different assignments to a set of k points $\mathbf{x}_1, \dots, \mathbf{x}_k$, we say these k points are *shattered* by \mathcal{F} .
 ▷ Sauer's lemma: $S_n(\mathcal{F}) \leq \sum_{i=0}^{d_{VC}} \binom{n}{i}$.
 * $S_n(\mathcal{F}) \begin{cases} = 2^n & n \leq d_{VC}; \\ \leq \left(\frac{d_{VC}e}{n}\right)^{d_{VC}} & n > d_{VC}. \end{cases}$
 ▷ If $d_{VC}(\mathcal{F}) < \infty$, then \mathcal{F} is PAC-learnable under 0-1 loss with sample complexity $\bar{n}(\epsilon, \delta) = C \cdot (d_{VC} + \log \frac{1}{\delta}) / (\epsilon^2)$ for some constant C . If $d_{VC} = \infty$, then \mathcal{F} is not PAC-learnable.

9 Unsupervised Learning

K-Means Clustering: Repeat the following 2 steps:

- ① Assign each point to the nearest cluster center:

$$\mathcal{D}_j = \{\mathbf{x} \in \mathcal{D} : j = \arg \min_{j'=1, \dots, K} \|\mathbf{x} - \mu_{j'}\|^2\}.$$

- ② Update cluster center to the average of points in that cluster:

$$\mu_j = \frac{1}{|\mathcal{D}_j|} \sum_{\mathbf{x} \in \mathcal{D}_j} \mathbf{x}.$$

- The objective $\sum_{j=1}^K \sum_{\mathbf{x} \in \mathcal{D}_j} \|\mathbf{x} - \mu_j\|^2$ is monotone non-increasing.

Maximum Likelihood Estimate:

$$\hat{\theta} = \arg \max_{\theta} \prod_{t=1}^n \Pr(\mathbf{x}_t; \theta) = \arg \max_{\theta} \sum_{t=1}^n \log \Pr(\mathbf{x}_t; \theta).$$

Appendix

Matrix Properties:

PSD	$\forall \mathbf{x} \in \mathbb{R}^n [\mathbf{x}^\top \mathbf{A} \mathbf{x} \geq 0]$	$\forall \lambda [\lambda \geq 0]$	\Leftrightarrow convex
PD	$\forall \mathbf{x} \neq \mathbf{0} [\mathbf{x}^\top \mathbf{A} \mathbf{x} > 0]$	$\forall \lambda [\lambda > 0]$	\Rightarrow strictly convex
NSD	$\forall \mathbf{x} \in \mathbb{R}^n [\mathbf{x}^\top \mathbf{A} \mathbf{x} \leq 0]$	$\forall \lambda [\lambda \leq 0]$	\Leftrightarrow concave
ND	$\forall \mathbf{x} \neq \mathbf{0} [\mathbf{x}^\top \mathbf{A} \mathbf{x} < 0]$	$\forall \lambda [\lambda < 0]$	\Rightarrow strictly concave
ID	none of the above	$\lambda_1 > 0; \lambda_2 < 0$	\Rightarrow neither nor

- $\mathbf{X}^\top \mathbf{X}$ is symmetric and PSD; $\mathbf{X}^\top \mathbf{X} + \lambda \mathbf{I}$ is PD.
- $\text{Eig}(\mathbf{A} + \mathbf{I}) = \text{Eig}(\mathbf{A}) + 1$. PSD + PD = PD.
- Product of eigenvalues is equal to determinant.
- Trace: ① linear ($\text{Tr}(\mathbb{E}[\mathbf{A}]) = \mathbb{E}[\text{Tr}(\mathbf{A})]$); ② $\mathbf{u}^\top \mathbf{v} = \text{Tr}(\mathbf{u}^\top \mathbf{v}) = \text{Tr}(\mathbf{v}^\top \mathbf{u})$.
- Derivative: $\nabla_{\mathbf{x}} \|\mathbf{A}\mathbf{x} + \mathbf{b}\|^2 = 2\mathbf{A}^\top (\mathbf{A}\mathbf{x} + \mathbf{b})$.